Introduction to Information Geometry

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https://franknielsen.github.io/IG/index.html

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What is information geometry? (1/4)

Consider a set of parametric probability distributions called the statistical model

For example, the set of normal distributions with mean μ and variance σ^2

$$\mathcal{P} = \left\{ p_{\lambda}(x) = \frac{1}{\sqrt{2\pi\sigma}} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right), \quad \lambda = (\mu,\sigma) \in \mathbb{H} \right\}$$

Parameter space Λ is the *upper plane* $\mathbb{H} = \mathbb{R} \times \mathbb{R}_{++}$



What kinds of geometric structures for this family of normal laws?

A few related questions:

- How to interpolate between two normals?
- How to define distances between them?
- Are they several ways to proceed?
 If so why? And how to choose the right structure?



What is information geometry? (2/4)

- Which **invariance principles** shall be satisfied by the geometric structures and the distances between statistical models
- First invariance principle: If we parameter Gaussians by (μ,σ²) or (μ,log(σ)) instead of (μ,σ), it should not change their distances nor the interpolating paths called ``geodesics''



Differential geometry of statistical models

- To each point of the manifold corresponds a unique parametric distribution:
- Statistical model is **identifiable** when $\lambda \leftrightarrow P_{\lambda}$
- Often a single global <u>chart</u> = atlas which covers the parameter domain



What is information geometry? (3/4)

Information geometry: study geometric structures on the manifold induced by identifiable statistical models

- Use language of geometry: geodesics, balls, information projection, statistical curvature and the tensor calculus. This tensor calculus made possible to study the efficiency of statistical estimators to higher order.
- Study the principles of invariance in statisticcs
- The new dual geometric structure can also be used beyond the statistical scope (pure geometry). For example, the information-geometric structures have been used to analyzed interior point methods in optimization
- Information geometry was born from the mathematical consideration of the Fisher metric and its induced geodesic distance to solve classification and statistical hypothesis test tasks

[Mahalanobis 1936] [Hotelling 1930] [Rao 1945]

Fisher information $I_{\chi}(\theta)$

 Consider a parametric family of laws indexed by D parameters Fisher Information Matrix (FIM) = covariance matrix of the score

$$X \sim p_{\theta}(x)$$
 $s_X(\theta) = \nabla_{\theta} \log p_{\theta}(x)$ $I_X(\theta) = \operatorname{Cov}[s_{\theta}]$

- FIM is symmetric and positive semi-definite (could be undefined too)
- FIM is said regular when positive-definite (yields the Fisher metric on manifolds)
- Interpreted as the **curvature** of the log-likelihood function:



Asymptotic normality of the MLE (Cramér-Rao lower bound)

For a general function f(x), the curvature is defined as

$$\kappa(x) = \frac{f''(x)}{(1 + (f'(x))^2)^{\frac{3}{2}}}, \quad r(x) = \frac{(1 + (f'(x))^2)^{\frac{3}{2}}}{|f''(x)|}$$

MLE: Maximum Likelihood Estimator

Radius of the **osculating circle** corresponds To the inverse of the absolute curvature

Fisher information $I_{\chi}(\theta)$

Fisher information small: Likelihood curvature is small (flat peak) Variance is large, low accuracy Fisher information large: Likelihood curvature is large (sharp peak) Variance is small, good accuracy

 $(\hat{\theta}_{\mathrm{MLE}}, l_x(\hat{\theta}_{\mathrm{MLE}}))$

Taylor 2nd-order with
$$l'(\hat{\theta}_{MLE}) = 0$$
: $l_x(\theta) \approx l_x(\hat{\theta}_{MLE}) + \frac{1}{2}(\theta - \hat{\theta}_{MLE})^2 l''(\hat{\theta}_{MLE}) - l''(\hat{\theta}_{MLE}) = -E_{\hat{\theta}_{MLE}}[l''(\theta)] = I(\hat{\theta}_{MLE}) \Rightarrow l_x(\theta) \approx l_x(\hat{\theta}_{MLE}) - \frac{1}{2}(\theta - \hat{\theta}_{MLE})^2 I(\hat{\theta}_{MLE})$

Two usual expressions of the Fisher information

• Using the first two **Bartlett identity** under the regularity condition that we can exchange k times the differentiation with the integration operations, we get

(Bartlett k) $\nabla^k E_{\theta}[\exp(l_x(\theta))] = E_{\theta} \left[\nabla^k \exp(l_x(\theta))\right] = \nabla^k E_{\theta}[p_{\theta}] = \nabla^k 1 = 0$

• Allows to rewrite the FIM under it is most famous forms:

$$I_{X}(\theta) = \operatorname{Cov}[s_{\theta}] \quad \operatorname{Cov}[s_{\theta}] = E[s_{\theta}s_{\theta}^{\perp}] - E[s_{\theta}]E[s_{\theta}]^{\perp}$$

$$(1) \quad First form: \qquad E[\nabla l_{x}(\theta)] = 0 \Rightarrow I_{X}(\theta) = E_{\theta} \left[\nabla \log p_{\theta}(x)(\nabla \log p_{\theta}(x))^{\top}\right]$$

$$(Bartlett k=1)$$

$$(2) \quad Second form (negative of the Hessian of the log-likelihood):$$

$$E_{\theta} \left[\nabla \log p_{\theta}(x)(\nabla \log p_{\theta}(x))^{\top}\right] + E \left[\nabla^{2} \log p_{\theta}(x)\right] = 0 \Rightarrow I_{X}(\theta) = -E[\nabla^{2} \log p_{\theta}(x)]$$

$$(Bartlett k=2)$$

$$I_{X}(\theta) = -E_{p_{\theta}} \left[\nabla^{2}l_{x}(\theta)\right] = -\left[\frac{\partial^{2}}{\partial \theta_{i}\partial \theta_{j}}l_{x}(\theta)l_{x}(\theta)\right]_{ij}$$

Mahalanobis and his generalized distance

- Motivated by the statistical analysis of human skulls collected in various regions. Each skull is characterized by d.
- Mahalanobis (1928, 1936) introduced the following D2 statistics and divergence entre deux groupes S₁ et S₂:

$$\Delta^2[p_{\mu_1,\Sigma}, p_{\mu_2,\Sigma}] = (\mu_1 - \mu_2)^\top \Sigma^{-1} (\mu_1 - \mu_2)$$

Precision matrix = inverse covariance

• Nowadays the **metric Mahalanobis distance**:

$$\Delta[p_{\mu_1,\Sigma}, p_{\mu_2,\Sigma}] = \sqrt{(\mu_1 - \mu_2)^{\top} \Sigma^{-1} (\mu_1 - \mu_2)}$$

Generalized the Euclidian distance when
$$\Sigma = I$$
, the identity matrix

Divergence = smooth dissimilarity which may be asymmetric and may not satisfy the triangular inequality of metric



P. C. Mahalanobis (1893-1972) Found of the Indian Statistical Institute (ISI)

Vol. VIII. } APRIL & SEPT. 1928. { Nos. 2 & 3.
I.—A STATISTICAL STUDY OF THE CHINESE HEAD
ВҰ
P. C. MAHALANOBIS

ON THE GENERALIZED DISTANCE IN STATISTICS.

By P. C. MAHALANOBIS.

(Read January 4, 1936.)

Mahalanobis distances: Vector spaces equipped with an inner product

Mahalanobis distance rewritten as

- $\Delta_{\Sigma}(\mu, \mu') = \|\mu \mu'\|_{\Sigma^{-1}} \\ \|x\|_{\Sigma^{-1}} = \sqrt{x^{\top}\Sigma^{-1}x}$
- For a symmetric positive-definite matrix (SPD) Q, we define the inner product by the following bilinear form: $\langle v_1, v_2 \rangle_O = v_1^{\mathsf{T}} Q v_2$
- Inner product induces a **norm** which in turn induces a metric **distance**:

$$\langle v_1, v_2 \rangle_E \to ||v||_E = \sqrt{\langle v, v \rangle} \to D_E(v_1, v_2) = ||v_1 - v_2||_E$$

 Inner product allows us to define the <u>orthogonality</u> between two vectors (and their subtended angle) and the vector <u>lengths</u>:

$$v_1 \perp v_2 \leftrightarrow \langle v_1, v_2 \rangle_Q = 0 \qquad ||v|| = \sqrt{\langle v, v \rangle}$$

• This geometry corresponds to the extrinsic geometry of tangent spaces of manifolds

Riemannian Fisher metric tensor field aka Fisher metric

• Consider the manifold $\mathcal{P} = \{p_{\theta}(x) : \theta \in \Theta\}$

 g_F : smooth fields of inner products on tangent planes

$$g_F(u,v) = [u]_B^{\mathsf{T}} I(\theta) [v]_B$$

vector **components**
$$[v]_B = (v^1, \dots, v^D)$$

expressed in the natural basis of the tangent plane Induced by the (local) chart $\theta(.)$ $\partial_i = \frac{\partial}{\partial \theta^i}$ $\partial_i = \frac{\partial}{\partial \theta^i}$

$$B = \{e_1 = \partial_1, \dots, e_D = \partial_D\}$$

$$g_{ij} = g_F(\partial_i, \partial_j) = I_{ij}(\theta)$$
$$g_F(u, v) = \sum_{i,j} g_{ij} u^i v^j = u^\top I(\theta) v$$

 $T_{p_{\theta}}$

 $q_F(u, v)$

 $T_{p'_{\theta}}\mathcal{P}$

Tangent plane representation for a manifold induced by a statistical model: Reinterpret the inner product

- On a tangent plane, we can choose any arbitrary basis to express vectors
- Inner product of two vectors is independent of the choice of basis: the component vectors depend on the basis but the vectors are geometric objects
- Express a vector v by a **representation** v(x)
- Basis vectors of T_{θ} can be chosen as the score vectors: $T_{\theta} = T_{p_{\theta}} = \{\sum_{i} v^{i} \partial_{i} l_{x}(\theta)\}$ $B = \{e_{1} = \partial_{1} l_{x}(\theta), \dots, e_{D} = \partial_{D} l_{x}(\theta)\}$ $e_{2} = \partial_{2} \log p_{\theta}(x)$
- The inner product can be reinterpreted as:

 $g_F(u,v) = E_{\theta}[u(x)v(x)] = \operatorname{Cov}(u(x), v(x))$

 $g_F(\partial_i, \partial_j) = E_\theta[\partial_i l_x(\theta)\partial_j l_x(\theta)]$

Expectation



Visualizing the Fisher metric and the Cramér–Rao bound

 $g_F(u,v) = [u]_B^\top I(\theta) [v]_B$

- Fisher metric:
- Visualize I(θ) by an ellipsoid
- Visualise the metric tensor field by **Tissot indicatrix**

Visualizing the Cramer-Rao lower bound :

- For each grid position (μ, σ) :
- Sample iid N (μ , σ)
- Maximum likelihood estimator of (μ, σ)
- Repeat k times to get an empirical estimator of the covariance matrix of the 2D parameters.
- Converge to the scaled inverse FIM

Inverse of the FIM displayed with Tissot indicatrix

projection

Mercator

$$\operatorname{Var}[\hat{\theta}_n] \succeq \frac{1}{n} I(\theta)^{-1}$$
$$I(\mu, \sigma) = \begin{bmatrix} \frac{1}{\sigma^2} & 0\\ 0 & \frac{2}{\sigma^2} \end{bmatrix}$$

 $\operatorname{Cov}[\hat{\theta}_n]$



 $\mathbb{H} = \mathbb{R} \times \mathbb{R}$

 \odot

Upper plane (μ,σ)

Rao distance on the Fisher-Rao manifold

$$D_{\text{Rao}}[p_{\theta_1}, p_{\theta_2}] = \rho_g(\theta_1, \theta_2) = \int_0^1 \sqrt{g_{\gamma(t)}(\dot{\gamma}(t), \dot{\gamma}(t))} \, \mathrm{d}t, \gamma(0) = \theta_1, \gamma(1) = \theta_2$$
$$= \int_0^1 ds_\theta(\gamma(t)) \, \mathrm{d}t \qquad \text{Here, } \gamma \text{ is the Riemannian geodesic}$$
(or add a minimizer on all paths γ)

 $D_{\mathrm{Rao}}[p_{\theta_1}, p_{\theta_2}] = \int_0^1 ds_{\theta}(\gamma(t)) \,\mathrm{d}t$

 $\mathrm{d}s_{\theta}(\gamma(t)$

Length element

$$ds_{\theta}^{2}(t) = \sum_{i=1}^{D} \sum_{j=1}^{D} g_{ij}(\theta) \dot{\theta}_{i}(t) \dot{\theta}_{j}(t)$$
$$\dot{\theta}_{k}(t) = \frac{d}{dt} \theta_{k}(t)$$

In practice:

- Need to calculated geodesics which are curves locally minimizing the length linking two endpoints (equivalently minimize the energy of squared length elements)
- Finding Fisher-Rao geodesics is a non-trivial tasks: No-known closed-form for the Fisher-Rao geodesic/distance between multivariate Gaussians!

Invariance under reparameterization of Rao's distance

Consider two different parameterizations of a statistical model:

$$\mathcal{P} = \{ p_\theta : \theta \in \Theta \} = \{ p_\eta : \eta \in H \}$$

Covariance transformation of the FIM under reparameterization

$$I_{\theta}(\theta) \xrightarrow{\eta = \eta(\theta)} I_{\eta}(\eta) = \left[\frac{\partial \theta_i}{\partial \eta_j}\right]^{\top} \times I_{\theta}(\theta(\eta)) \times \left[\frac{\partial \theta_i}{\partial \eta_j}\right]$$

... However the length element is **invariant** : $ds_{\theta} = ds_{\eta}$

So that the Fisher-Rao distance is **invariant** :

$$\rho_{\text{Rao}}(p_{\eta_1}, p_{\eta_2}) = \rho_{\text{Rao}}(p_{\theta_1}, p_{\theta_2})$$

> This is the first principle of invariance of information geometry

Fisher-Rao geometry of univariate normal distributions



In general, location-scale families yield a hyperbolic Fisher-Rao geometry

Fisher-Rao manifolds: Intrinsic vs extrinsic viewpoints

• A Riemannian manifold of dimension D can be embedded as a surface of Euclidian space in dimension O(D²) :

Isometric embedding of the manifold

• For example, Rao's distance between two Bernoulli distributions or categorial distributions can be easily found by embedding the standard simplex on the positive orthant of the 1D sphere of radius 2 in R2 by the 2 x square root transformation Covariance transformation of the FIM



Neuromanifolds and deep learning

- A neural network (like the multilayer perceptrons, MLPs) is described by a feed-forward architecteur by means of a large number of parameters θ
- Consider stochastic neural networks (SNNs) with noisy output: $y = NN_{\theta}(x) + \epsilon$

... Eg a Gaussian noise:
$$p_{\theta}(x,y) = p(x)p_{\theta}(y|x) = \frac{p(x)}{\sqrt{2\pi}\sigma} \exp\left(-\frac{1}{2\sigma^2}(y - NN_{\theta}(x))^2\right)$$

• Neuromanifold is

$$\mathcal{P} = \{ p_{\theta}(x, y) : \theta \in \Theta \}$$

Modelled by a neuromanifold Me

 \boldsymbol{h}_{L-1}

 h_1

Learning trajectory

Parameter space θ :

 $\boldsymbol{\theta}_L$

 $\boldsymbol{\theta}_1$

Artificial

Neural Network

SN 2017

- Maximizing the likelihood of a SNN with Gaussian noise amounts to minimize the mean quadratic error
- Given a training set, we learn the parameter of the NNs using gradient descent. We can visualize the learning process as a trajectory on the neuromanifold modelling the parameter space. We observe plateau phenomena when nearning a singularity on the manifold where the Fisher information matrix is rank deficient or close to (small eigenvalues of the FIM).

Natural gradient: Steepest Riemannian descent

Ordinary gradient descent (GD) method for minimizing a loss function E(.) :

$$heta_{t+1} = heta_t - lpha
abla E(heta_t)$$
Learning step or rate

- Ordinary GD depends on the parameterization
- Plateau phenomena near singularities (almost degenerate Fisher information)

Natural gradient is invariant to reparameterization and avoids plateaus:

$$\tilde{\nabla}E(\theta) := G(\theta)^{-1} \nabla_{\theta} E(\theta)$$

$$\tilde{\nabla}E_{\eta}(\eta) = \tilde{\nabla}E_{\theta}(\theta)$$

Natural gradient descent (NGD)

$$\theta_{t+1} = \theta_t - \alpha \tilde{\nabla} E(\theta_t)$$

Where α = step size

Natural gradient descent is different from the **Riemannien gradient descent** which relies on the Riemannian exponential map which is time consuming (retraction)₁₉

What is information geometry? (4/4)

- A dual structure which allows to explain the duality between statistical inference like the maximum likelihood estimator and a family of statistical models obtained from the maximum entropy principle: Information geometry explains the link between Shannon entropy, the Kullback-Leibler divergence and exponential families in statistics.
- Second principle of invariance by sufficient statistic

This core dual structure of information geometry:

- Open new perspectives: For example, non-extensive entropies like Tsallis entropy, complex systems, conformal geometry of deformed exponential families, etc.
- Many applications of information geometry ranging from signal processing (Radar, Brain-Machine interfances, etc.), to medical imaging, to machine learning and AI, etc.

 $p_{\lambda}(x)$

 $x \in \text{sample space } \Omega$

Geodesics are defined according to affine connections

- In Riemannian geometry, geodesics are locally length minimizing curves
- Geodesics $\gamma(t)$ defined by connection ∇ as ∇ -autoparallel curves:

$$\nabla_{\dot{\gamma}}\dot{\gamma} = 0, \quad \dot{\gamma} = \frac{d}{dt}\gamma(t) \quad \frac{d^2\theta_k}{dt^2} + \sum_{i=1}^p \sum_{j=1}^p \Gamma_{ij}^k \frac{d\theta_i}{dt} \frac{d\theta_j}{dt} = 0, \quad k = 1, \dots, p,$$

where $\nabla_X T$ is the <u>covariant differentiation operator and X</u> is a vector field D³ Christoffel symbols Γ which are functions characterizing the affine connection ∇ (covariant derivative)

• In Riemannian geometry, we use by default the <u>Levi-Civita connection</u> which is derived from the metric tensor field g (thus implicit in Rie. Geo.) :

$$\Gamma_{ij}^{k} = \frac{1}{2} \sum_{m=1}^{p} \left(\frac{\partial g_{im}(\boldsymbol{\theta})}{\partial \theta_{j}} + \frac{\partial g_{jm}(\boldsymbol{\theta})}{\partial \theta_{i}} - \frac{\partial g_{ij}(\boldsymbol{\theta})}{\partial \theta_{m}} \right) g^{mk}(\boldsymbol{\theta}), \quad i, j, k = 1, \dots, p,$$

Affine connection ∇ : Visualizing the curvature by the ∇ -parallel transport along smooth loops







Élie Cartan 1869-1951

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A connection is flat is there exists a coordinate system θ for which the Christoffel symbols all vanish: Γ(θ)=0 \rightarrow Geodesics are plotted as line segments in the local chart θ

$$\frac{d^2\theta_k}{dt^2} + \sum_{i=1}^p \sum_{j=1}^p \mathbb{D}_{ij}^k \frac{d\theta_i}{dt} \frac{d\theta_j}{dt} = 0, \quad k = 1, \dots, p,$$

Geodesics of flat connection

Line segments

The key dual structure of information geometry

$$(M,g,
abla,
abla^*)$$
 such that

$$\frac{\nabla + \nabla^*}{2} = {}^g \nabla$$

Given a torsion-free affine connection ∇ and a metric tensor g, we can build a unique dual torsion free connection ∇* such that the metric is preserved by the bi-parallel transport

$$\langle u, v \rangle_{c(0)} = \left\langle \prod_{c(0) \to c(t)}^{\nabla} u, \prod_{c(0) \to c(t)}^{\nabla^*} v \right\rangle_{c(t)} \cdot \left\langle v_1 \right\rangle_{c(t)} \cdot \left\langle v_1$$

- The dual connection of the dual connection is the original connection $(\nabla^*)^* = \nabla$.
- Question: How to find meaningful dual connections?
 - Method of Amari-Nagaoka (1982) : the statistical expected α -connexions (Chentsov 1972)
 - Method of Eguchi (1983): Build dual connections from dual divergences (contrast functions) 23

The dual α -geometry of Amari and Nagaoka

Structure $(\mathcal{P}, g_F, \nabla^{lpha}, \nabla^{-lpha})$ Dual connections with respect to Fisher metric

 ∇^{α} Defined by the Christoffel symbols $\Gamma^{\alpha}_{ij,k} = E_{\theta} \left[\left(\partial_i \partial_j l + \frac{1-\alpha}{2} \partial_i l \partial_j l \right) \partial_k l \right]$

Some α -connections:

- O-connection = Levi-Civita metric connection of Fisher metric : Fisher-Rao mfd
- 1-connection is called the exponential connection
- -1 connection is called the mixture connection

[Efron 1975] [Dawid 1975]

$$(\nabla^e)^* = \nabla^m \quad (\nabla^m)^* = \nabla^e$$

 $\nabla^\alpha = \frac{1+\alpha}{2} \nabla^e + \frac{1-\alpha}{2} \nabla^m$

Dual geometry em used to study the duality between estimators/stat models

Eguchi's dual geometry induced by a divergence

• Structure $(M, {}^Dg, {}^D\nabla, {}^D\nabla^*)$

Divergence information geometry

(self dual when divergence is symmetric)

• Get a **divergence** (contrast function) from a statistical divergence between parametric distributions. For example, the Kullback-Leibler divergence between two parametric distributions from a family P:

Parametric divergence contrast divergence

$$D_{\mathrm{KL}}^{\mathcal{P}}(\theta_1:\theta_2) := D_{\mathrm{KL}}[p_{\theta_1}:p_{\theta_2}]$$

 $\equiv {}^{D^{*}}\nabla$

Statistical divergence

- Eguchi Levi-Civita metric associated to D is ${}^{D}g_{ij}(\theta_1) = -\partial_{\theta_1}\partial_{\theta_2}D(\theta_1:\theta_2)|_{\theta_1=\theta_2}$
- Eguchi connection associated to D ${}^{D}\Gamma_{ij,k} = g\left({}^{D}\nabla_{\partial_{i}}\partial_{j},\partial_{k}\right) = -\partial_{\theta_{1}^{i}}\partial_{\theta_{2}^{k}}D\left(\theta_{1}\|\theta_{2}\right)\Big|_{\theta_{1}=\theta_{2}}$
- Define the dual divergence en by swapping the parameter order:

$$D^*(\theta_1:\theta_2) := D(\theta_2:\theta_1)$$

 $\int^{D_g} \nabla = \frac{1}{2}$

• Get dual affine connections

Levi-Civita connection is recovered from

f-divergences and their induced connections

• Relative entropy or the Kullback-Leibler divergence belongs to a broader class of dissimilarities : **f-divergences** [Csiszar'63] [Ali&Silvey'66]

$$I_{f}[p:q] = \int pf(q/p)d\mu \longrightarrow D_{\mathrm{KL}}[p:q] = \int p\log p/qd\mu = I_{f_{\mathrm{KL}}}[p:q] \quad f_{\mathrm{KL}}(u) = -\log u$$

• Generator f(.) is convex, strictly convex at 1.

- WLOG, fix f'(1)=0 et f''(1)=1 to get a **standard f-divergence**
- Dual f-divergence $I_f^*[p:q] = I_f[q:p] = I_{f^*}[p:q]$ with $f^*(u) = uf(1/u)$
- The Eguchi induced metric tensor of std f-divergences = Fisher : $I_f^{\mathcal{P}} g = g_F = I_{f^*}^{\mathcal{P}} g$
- Induced f-connections wrt to f-divergences between distributions of a family P match with the α -connections of Amari and Nagaoka :

$${}^{I_f}\mathcal{P} = \mathcal{P} \nabla^{\alpha_f} \qquad \qquad \alpha_f = 3 + 2 \frac{f'''(1)}{f''(1)}$$

Statistical distances and information monotonicity

 Consider a transformation Y=t(X) on random variables between two measurable spaces (deterministic or stochastic, Markov kernel):

$$t: (\mathcal{X}, \Sigma) \to (\mathcal{Y}, \Sigma') \qquad Y_i = t(X_i)$$

• Second principle of invariance: We should not increase the power of discrimination of divergences by a transformation: $/ / \sqrt{2} = \sqrt{2} \sqrt{2} \sqrt{2}$

$$D[p_{X_1}:p_{X_2}] \ge D[q_{Y_1}:q_{Y_2}] \ge 0$$

- Fisher information monotonicity: $I_{t(X)}(\theta) \leq I_X(\theta)$
- Equality holds if and only if t(X) is a sufficient statistic
- A sufficient statistic summarizes all necessary information for inference on the parameter heta (statistical lossless compression): $\Pr(x| heta) = \Pr(x|t)$
- Theorem: f-divergences are the only separable monotone divergences

 $D[p_{X_1}:p_{X_2}] \ge D[q_{Y_1}:q_{Y_2}] \ge 0$

Exponential families have finite dim. sufficient statistic vector

• An exponential family is a set of parametric distributions with density which can be expressed canonically as : μ (eg., Lebesgue or counting measure)

 $p_{\theta}(x) = \exp(\langle \theta, t(x) \rangle - F(\theta))$ By default, scalar product = Euclidean inner product where F is an analytic and stricty convexe and differentiable function: $F(\theta) = \log \int \exp(\theta x) d\mu(x)$ F: log partition function or cumulant function

Natural parameter space for full EFs $\Theta = ig\{ heta \ : \ \int \exp(heta x) \mathrm{d} \mu(x) < \infty ig\}$



Information geometry of exponential families: Dually flat

- Statistical model: natural exp. family $\mathcal{P} = \{p_{\theta}(x) = \exp(x^{\top}\theta F(\theta))\}$ or more generally $\mathcal{P} = \{\exp(\theta^{\top}t(x) - F(\theta))h(x)\}$ $p_{\theta}(x) = p^{\eta}(x)$
- Exponential connection and dual mixture connection are both flat:

Dually flat spaces of exponential families : $(M, g_F, \nabla^e, \nabla^m)$

- Fisher information metric is a Hessian metric $g_F(\theta) = I(\theta) = \text{Cov}[t(X)] = \nabla^2 F(\theta)$
- By using the Legendre-Fenchel transformation, we get a dual coordinate system eta $F^*(\eta) = \sup \theta^\top \eta F(\theta), \eta = \nabla F(\theta) = E[t(X)]$
- Moment or mean parameterization:

$$p_{\theta}(x) = p^{\eta}(x)$$



• Fisher information matrix can be expressed in the moment parameter: $I(\eta) = \nabla^2 F^*(\eta) = g_F(\theta)^{-1}$

Dually flat spaces with Hessian structures

 The primal and dual geodesics are line segments in the affine theta and eta coordinate system:

$$p_{\theta}(x) = p^{\eta}(x)$$

Primal geodesic

$$\gamma_{p_{\theta_1}p_{\theta_2}}(t) = p_{(1-t)\theta_1 + t\theta_2}$$

Dual geodesic

$$\gamma_{p_{\theta_1} p_{\theta_2}}^*(t) = p^{(1-t)\eta_1 + t\eta_2}$$





 In dually flat spaces, there is a canonical divergence: The Bregman divergence. For exponential families, this Bregman divergence amounts to the dual divergence of the Kullback-Leibler divergence (=reverse KLD) between corresponding densities :

$$B_F(\theta_1:\theta_2) = D_{\mathrm{KL}}^*[p_{\theta_1}:p_{\theta_2}] = D_{\mathrm{KL}}[p_{\theta_2}:p_{\theta_1}]$$

Visualizing a Bregman divergence as a vertical gap

- Let $F(\theta)$ be a strictly convex and differentiable function defined on an open convex domain Θ
- Bregman divergence interpreted as the vertical gap between

point $(\theta_1, F(\theta_1))$ and the linear approximation of $F(\theta)$ at θ_2 evaluated at θ_1 :



Mixed coordinates and the Legendre-Fenchel divergence

- Dual <u>Legendre-type</u> functions
- Convex conjugate of F is
- Fenchel-Young inequality :

$$\theta = \nabla F^*(\eta) \qquad \qquad \eta = \nabla F(\theta)$$
$$F^*(\eta) = \eta^\top \nabla F^*(\eta) - F(\nabla F^*(\eta))$$

$$F(\theta_1) + F^*(\eta_2) \ge \theta_1^\top \eta_2$$

 $abla F^* = (
abla F)^{-1}$ Gradient
are inverse
of each other

with equality holding if and only if $\eta_2 = \nabla F(\theta_1)$

• Fenchel-Young divergence make use of the mixed coordinate systems θ et η to express a Bregman divergence as $B_F(\theta_1 : \theta_2) = Y_{F,F^*}(\theta_1 : \eta_2)$

$$Y_{F,F^*}(\theta_1:\eta_2) := F(\theta_1) + F^*(\eta_2) - \theta_1^\top \eta_2 = Y_{F^*,F}(\eta_2,\theta_1)$$

Dual Bregman and dual Fenchel-Young divergences

• Identity for dual Bregman divergences:

 $B_F(\theta_1:\theta_2) = B_{F^*}(\eta_2:\eta_1)$

(The Bregman divergence coincides with the reverse Bregman divergence for the convex dual generator)

- By definition, dual divergence = divergence on swapped parameter order: $D^*(\theta_1 : \theta_2) := D(\theta_2 : \theta_1)$
- Thus in a dually flat space, we can write the canonical divergence as :

$$B_F(\theta_1:\theta_2) = Y_{F,F^*}(\theta_1:\eta_2) = Y_{F^*,F}(\eta_2,\theta_1) = B_{F^*}(\eta_2:\eta_1)$$

On a Bregman manifold, we can thus get 2ⁿ equivalent formula with n terms

Generalized Pythagoras theorem in dually flat spaces

Generalized Pythagoras' theorem Pythagoras' theorem in the Euclidian geometry orthogonality condition: **Self-dual** $(\eta(p) - \eta(q))^{\top}(\theta(r) - \theta(q)) = 0$ $F_{\text{Eucl}}(\theta) = \frac{1}{2} \theta^{\top} \theta$ $g_{F_{\text{Euc}}} = I$ $\gamma_{pq}(t)$ $B_{F_{\text{Eucl}}}(\theta_1:\theta_2) = \frac{1}{2}\rho_{\text{Eucl}}^2(\theta_1,\theta_2)$ С $\gamma_{pq} \perp_q \gamma_{qr}^*$ b $\gamma_{ar}^*(t')$ а

 $D_F(\gamma_{pq}(t):\gamma_{qr}(t')) = D_F(\gamma_{pq}(t):q) + D_F(q:\gamma_{qr}^*(t')), \quad \forall t, t' \in (0,1).$

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 $a^{2} + b^{2} = c^{2}$

Identity of Bregman divergence with three parameters

 $B_F(\theta_1:\theta_2) = B_F(\theta_1:\theta_3) + B_F(\theta_3:\theta_2) - (\theta_1 - \theta_3)^\top (\nabla F(\theta_2) - \nabla F(\theta_3)) \ge 0$

Information projection uniqueness theorems

- Define the e-projection and the m-projection of a point onto a submanifold with respect to the affine connections ∇^e (∇⁺¹) and ∇^m (∇⁻¹) where the orthogonality is given by the Fisher information g_F
- A submanifold is e-flat if and only if when expressed in the θ-coordinate system, we get an affine subspace.
- Similar definition for a $\, \textbf{m-flat}$ submanifold wrt η
- Generalized Pythagoras' theorem allows to prove that the e-projection of a point onto a n-flat submanifold is <u>unique</u> and corresponds the minimization of a Bregman divergence. Similarly, m-projection of a point onto a e-flat submanifold is unique and can be obtained as the minimization wrt to the dual Bregman divergence





Maximum likelihood estimator as a m-projection

Let $\{x_1, ..., x_n\}$ be i.i.d variates of an **exponential family** P (e-flat) The empirical distribution is called the <u>observed point</u>



By considering an arbitrary divergence D[.:.] instead of the Kullback-Leibler divergence we get **D-estimators**. MLE is interpreted as the KLD-estimator

Maximum entropy and e-projection

• Given observations with $E[t_i(x)]=m_i$, the maximum entropy principle of Jaynes estimate the distribution which maximizes Shannon entropy under the moment contraints $E_p[t_1(X)] = m_1, \dots, E_p[t_D(X)] = m_D$.



- Set of distributions maximizing entropy under the constraints $E[t(x)]=\eta$ for all η form an exponential family $p^* \in \mathcal{E} := \{p_{\theta}(x) = \exp(\sum t_i(x)\theta_i F(\theta))\}$
- For example, the MaxEnt distributions for $E[x]=\eta_1$ et $E[x^2]=\eta_2$ yield the family of normal distributions (univariate of order 2, dim. of natural parameter space)

Alternating projections: The em algorithm (= $\nabla^{e} \nabla^{m}$) Find the minimal distance between two submanifolds $\min_{p \in \mathcal{P}} \min_{q \in \mathcal{Q}} D_{\mathrm{KL}}[p:q]$ = Solve jointly the following minimization: • When a submanifold P is m-flat and the submanifold Q $D_{\mathrm{KL}}[\mathcal{P}:\mathcal{Q}] = D_{\mathrm{KL}}[p^*:q^*]$ is e-flat then we get unique sequence of e/m *m*-flat submanifold \mathcal{P} alternating projections. Starting from q₁ we repeat: p_{t+1} n^* e-projection : $p_{t+1} = \arg\min_{p \in \mathcal{P}} D_{\mathrm{KL}}[p:q_t]$ m-projection : $q_{t+2} = \arg\min_{q \in \mathcal{Q}} D_{\mathrm{KL}}[p_{t+1}:q]$ q_{t+2} q_t q0 e-flat sub-manifold QThe em algorithm is useful for : converge towards the pair of points which minimize - Interpreting the EM alg. in statistics - To analyze generative models in the Kullback-Leibler divergence between P and Q : deep learning like the VAEs or GANs

$$D_{\mathrm{KL}}[p^*:q^*] = \min_{p \in \mathcal{P}} \min_{q \in \mathcal{Q}} D_{\mathrm{KL}}[p:q] = \lim_{t \to \infty} D_{\mathrm{KL}}[p_{t+1}:q_t]$$

Bregman cyclic projections (in a chart)

• Let n convex objects O₁, ..., O_n be defined in a

 θ -coordinate system on a convex Θ

- Goal: find a common point in the intersection of these objects if intersection is non-void
- Repeat <u>cyclically</u> the **Bregman projections:**

 $\theta_0 \in \Theta, t \leftarrow 0$

$$\theta_{t+1} = \arg\min_{\theta \in O_{1+(t \mod n)}} B_F(\theta_t : \theta)$$

• This sequence converges towards a common point for non-empty intersection



Chernoff information and Bayesian hypothesis tests

- Let P_1 and P_2 be two distributions, and take n i. i. d. variates $x_1, ..., x_n$ from the statistical mixture model 1/2 P_1 + 1/2 P_2
- Which rule to classify these n samples with labels P1 or P2?
- Best rule minimizing the probability of error is maximum a posteriori (MAP)
- Probability of error is bounded by $P_e^n = 2^{-nC(P_1,P_2)}$ where C is the following Chernoff divergence (or Chernoff information)

$$C(P,Q) = -\log\min_{\alpha \in (0,1)} \int p^{\alpha}(x)q^{1-\alpha}(x)d\nu(x).$$

When P_1 and P_2 are two densities of a same exponential family, we have:

$$C(P_{\theta_1} : P_{\theta_2}) = B(\theta_1 : \theta_{12}^{(\alpha^*)}) = B(\theta_2 : \theta_{12}^{(\alpha^*)})$$

where α^* is the optimal exponent in (0,1)

$$P^* = P_{\theta_{12}^*} = G_e(P_1, P_2) \cap \operatorname{Bi}_m(P_1, P_2)$$

$$(p_{-12}) \qquad (p_{-12}) \qquad (p_{-12$$



- Voronoi diagram wrt
- Kullback-Leibler divergence

Exponential family manifold (dually flat) Bregman Voronoi diagram

Computational geometry to calculate the Bregman Voronoi diagrams

Natural gradient in dually flat spaces

On a global Hessian manifold (Bregman manifold induced by a convex function F), the Fisher information matrix can be expressed

$$I_{\theta}(\theta) = \nabla_{\theta}^{2} F(\theta) = \nabla_{\theta} \nabla_{\theta} F(\theta) = \nabla_{\theta} \eta$$

Définition du paramètre moment
$$\tilde{\nabla}_{\theta} I_{\theta}(\theta) := I^{-1}(\theta) \nabla_{\theta} I_{\theta}(\theta)$$

Find many applications in optimization in machine learning: Natural evolution strategies (NES), Bayesian inference, etc.

To summarize information geometry in 1 slide!

- Geometric structures for a parametric family of distributions: the statistical model
- Invariance wrt distribution parameterizations

 (θ) and sufficient statistics
 (on sample space Ω).
 Distance cannot increase by a measurable
 transformation Y=t(X), and does not change only
 if t is a sufficient transformation
- Fisher-Rao geometry equipped with the Rao Riemannian geodesic length distance
- Dual α-geometry (they are not necessarily associated divergences, except when dually flat)
- Interpret statistical estimator (maximum likelihood estimator) and statistical model (maximum entropy): Pythagoras' theorem and information projections in dually flat spaces (e.g., exponential/mixture families)



Fisher-Rao Riemannian geometry



Thank you (1/2)



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